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#### Formalization: Preference Systems

**Notation:** For a preorder R, denote by  $P_R$  its *strict* and by  $I_R$  its *indifference* part.

**Definition 1.** Let  $A \neq \emptyset$  be a set. Let  $R_1 \subseteq A \times A$  be a preorder on A and  $R_2 \subseteq A$  $R_1 \times R_1$  be a preorder on  $R_1$ . The triplet  $\mathcal{A} = [A, R_1, R_2]$  is called a **preference** system on A. We call  $\mathcal{A}$  consistent if  $\exists u : A \rightarrow [0,1]$  s.t. for all  $a, b, c, d \in A$ :

#### Spaces with Differently Scaled Dimensions

Consider an r-dimensional space  $A \subseteq \mathbb{R}^r$  and assume that

- the first  $0 \le z \le r$  dimensions are of cardinal scale and
- the remaining dimensions are purely ordinal.

Utilize the cardinal information only on parts of A where there is no possible conflict

- $(a,b) \in R_1 \Rightarrow u(a) \ge u(b)$  (with  $= iff \in I_{R_1}$ ).
- $((a,b), (c,d)) \in R_2 \Rightarrow u(a) u(b) \ge u(c) u(d)$  (with = iff  $\in I_{R_2}$ ). The set of all representations u of  $\mathcal{A}$  is denoted by  $\mathcal{U}_{\mathcal{A}}$ .

**Definition 2.** A consistent preference system  $\mathcal{A}$  is **bounded**, if  $\exists a_*, a^* \in A$  such that  $(a^*, a) \in R_1$ , and  $(a, a_*) \in R_1$  for all  $a \in A$ , and  $(a^*, a_*) \in P_{R_1}$ . In this case, for  $\delta \in [0,1)$ , denote by  $\mathcal{N}^{\delta}_{A}$  the set of all  $u \in \mathcal{U}_{A}$  with  $u(a_{*}) = 0$ ,  $u(a^{*}) = 1$ , and  $u(a) - u(b) \ge \delta \quad \land \quad u(c) - u(d) - u(e) + u(f) \ge \delta$ 

for all  $(a, b) \in P_{R_1}$  and for all  $((c, d), (e, f)) \in P_{R_2}$ .

#### SITUATION GENERAL

Generalized Stochastic Dominance (GSD)

For  $\pi$  a probability measure on  $(\Omega, \mathcal{S})$  and  $\mathcal{A}$  a consistent preference system, set  $\mathcal{F}_{(\mathcal{A},\pi)} := \Big\{ X \in A^{\Omega} : u \circ X \in \mathcal{L}^1(\Omega, \mathcal{S}, \pi) \; \forall u \in \mathcal{U}_{\mathcal{A}} \Big\}.$ For  $X, Y \in \mathcal{F}_{(\mathcal{A},\pi)}$ , say Y is  $(\mathcal{A},\pi)$ -dominated by X, formally  $(X,Y) \in R_{(\mathcal{A},\pi)}$ , if  $\forall u \in \mathcal{U}_A : \mathbb{E}_{\pi}(u \circ X) > \mathbb{E}_{\pi}(u \circ Y).$ 

The preorder  $R_{(\mathcal{A},\pi)}$  on  $\mathcal{F}_{(\mathcal{A},\pi)}$  called **generalized stochastic dominance (GSD)**.

# GENERAL & SITUATION

Testing for GSD

with the ordinal one. Consider A to be a subsystem of  $pref(\mathbb{R}^r) = [\mathbb{R}^r, R_1^*, R_2^*]$ , where  $R_1^* = \left\{ (x, y) : x_j \ge y_j \; \forall j \le r \right\}$  $R_{2}^{*} = \left\{ ((x, y), (x', y')) : \begin{array}{c} x_{j} - y_{j} \ge x'_{j} - y'_{j} \quad \forall j \le z \\ x_{j} \ge x'_{j} \ge y'_{j} \ge y_{j} \quad \forall j > z \end{array} \right\}.$ 

CONCRETE & SDSD

## **Example:** Poverty Analysis

We use the ALLBUS data and account for three dimensions of poverty: income (numeric), health (ordinal, 6 levels) and education (ordinal, 8 levels). E.g., for the following two pairs of vectors we can utilize the cardinal dimensions:



Some Properties of SDSDs

Assume *i.i.d.* samples  $\mathbf{X} = (X_1, \ldots, X_n)$  and  $\mathbf{Y} = (Y_1, \ldots, Y_m)$  of X and Y.

Hypotheses:

 $H_0: (Y, X) \in R_{(\mathcal{A}, \pi)}$  vs.  $H_1: (Y, X) \notin R_{(\mathcal{A}, \pi)}$ 

Test Statistic:

$$\omega \mapsto \inf_{u \in \mathcal{N}_{\mathcal{A}_{\omega}}^{\delta_{\varepsilon}(\omega)}} \sum_{z \in (\mathbf{X}\mathbf{Y})_{\omega}} u(z) \cdot (\hat{\pi}_{X}^{\omega}(\{z\}) - \hat{\pi}_{Y}^{\omega}(\{z\}))$$

with, for  $\omega \in \Omega$  and  $\varepsilon \in [0, 1]$  fixed, and

- $\hat{\pi}_X^{\omega}$  and  $\hat{\pi}_Y^{\omega}$  the observed empirical image measures of X and Y,
- $(\mathbf{XY})_{\omega} = \{X_i(\omega) : i \le n\} \cup \{Y_i(\omega) : i \le m\} \cup \{a_*, a^*\}, \text{ and }$
- $\mathcal{A}_{\omega}$  the subsystem of  $\mathcal{A}$  restricted to  $(\mathbf{X}\mathbf{Y})_{\omega}$ , and
- $\delta_{\varepsilon}(\omega) := \varepsilon \cdot \sup\{\xi : \mathcal{N}_{\mathcal{A}}^{\xi} \neq \emptyset\}.$

**Computation:**  $d_{\mathbf{X},\mathbf{Y}}^{\varepsilon}$  can be computed by solving one single *linear program*. **Test scheme:** We made observations of the i.i.d. variables, i.e., we observed:  $\mathbf{x} := (x_1, \dots, x_n) := (X_1(\omega_0), \dots, X_n(\omega_0))$ ,  $\mathbf{y} := (y_1, \dots, y_m) := (Y_1(\omega_0), \dots, Y_m(\omega_0))$ As the worst case of  $H_0$  is  $\pi_X = \pi_Y$ , we can perform a *permutation test*: **Step 1:** Pool data sample:  $\mathbf{w} := (w_1, \ldots, w_{n+m}) := (x_1, \ldots, x_n, y_1, \ldots, y_m)$ 

**Theorem 1.** Let  $X = (\Delta_1, \ldots, \Delta_r), Y = (\Lambda_1, \ldots, \Lambda_r) \in \mathcal{F}_{(pref(\mathbb{R}^r), \pi)}$ . Then:

 $pref(\mathbb{R}^r)$  is consistent. i)

FOR SDSDs

- If z = 0, then  $R_{(pref(\mathbb{R}^r),\pi)}$  equals (first-order) stochastic dominance w.r.t.  $\pi$ ii) and  $R_1^*$  (short:  $FSD(R_1^*, \pi)$ ).
- If  $(X, Y) \in R_{(pref(\mathbb{R}^r),\pi)}$  and  $\Delta_j, \Lambda_j \in \mathcal{L}^1(\Omega, \mathcal{S}_1, \pi)$  for all  $j = 1, \ldots, r$ , then iii) I.  $\mathbb{E}_{\pi}(\Delta_j) \geq \mathbb{E}_{\pi}(\Lambda_j)$  for all  $j = 1, \ldots, r$ , and
  - II.  $(\Delta_j, \Lambda_j) \in FSD(\geq, \pi)$  for all  $j = z + 1, \ldots, r$ .

If all components of X are jointly independent and all components of Yare jointly independent, I. and II. imply  $(X, Y) \in R_{(pref(\mathbb{R}^r), \pi)}$ .

### Test in the Example

For the ALLBUS data, we focus on subsamples with n = m = 100 men and women.



**Step 2:** For all  $I \subseteq \{1, \ldots, n+m\}$  with |I| = n, compute  $d_{\mathbf{X}\mathbf{Y}}^{\varepsilon}$  for  $(w_i)_{i\in I}$  and  $(w_i)_{i \in \{1,\dots,n+m\}\setminus I}$  instead of  $\mathbf{x}/\mathbf{y}$  to get  $d_I^{\varepsilon}$ . Sort all  $d_I^{\varepsilon}$  increasingly to get  $d_{(1)}^{\varepsilon},\dots,d_{(k)}^{\varepsilon}$ . **Step 3:** Reject  $H_0$  if  $d_{\mathbf{X},\mathbf{Y}}^{\varepsilon}(\omega_0) > d_{(\ell)}^{\varepsilon}$ , with  $\ell := \lceil (1-\alpha)k \rceil$  and  $\alpha$  the significance level.

# GENERAL & SITUATION Robustifying the Test

**Idea:** Use *credal sets* to robustify the permutation test. Concretely, allow the samples to be (potentially) *biased* in the sense that we only assume the *true empirical laws* to lie in some credal neighborhoods  $\mathcal{M}_X$  and  $\mathcal{M}_Y$  around the biased empirical laws. Adapted test scheme: Replace

- $d_{\mathbf{X},\mathbf{Y}}^{\varepsilon}(\omega_0)$  by  $\inf_{(\pi_1,\pi_2)\in\mathcal{M}_X^{\omega_0}\times\mathcal{M}_Y^{\omega_0}} d_{\mathbf{X},\mathbf{Y}}^{\varepsilon}(\omega_0)$
- $d_I^{\varepsilon}(\omega_0)$  by  $\sup_{(\pi_1,\pi_2)\in\mathcal{M}_X^{\omega_0}\times\mathcal{M}_Y^{\omega_0}} \tilde{d}_I^{\varepsilon}(\omega_0)$

**Results in:** Valid (yet conservative) statistical test!

**Results:** All tests significant for  $\alpha = 0.05$ . P-values decrease with increasing regularization strength  $\varepsilon$  of the test statistic.

## Credal Sets in Example

A special class of credal sets are  $\gamma$ contamination models. For  $\omega \in \Omega$ ,  $\gamma \in [0, 1]$ , and  $Z \in \{X, Y\}$ , we set  $\mathcal{M}_{Z}^{\omega} = \Big\{ \pi : \pi \ge (1 - \gamma) \cdot \hat{\pi}_{Z}^{\omega} \Big\}.$ **Interpretation:** The contamination parameter  $\gamma$  can be interpreted as the share of data that can deviate from the *i.i.d.* sampling assumption.

#### **Results Robust Testing**

